

Probability Density Functions of Sums, Products, Quotients, and Powers of Random Variables (PDFs of $X + Y$, XY , X/Y , and X^Y)

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X and Y are mutually independent continuous random variables, and their probability density functions (PDFs) are given by $f(x)$ and $g(y)$, respectively. We derive the probability density functions of (a) $Z = X + Y$, (b) $Z = XY$, (c) $Z = X/Y$, and (d) $Z = X^Y$.

(a) $Z = X + Y$

$$\begin{aligned} p(z) &= \int \delta(z - x - y) f(x) g(y) dx dy \\ &= \int f(z - y) g(y) dy \end{aligned} \tag{1}$$

【Example 1】 When the random variables x and y follow the uniform distribution on $0 \leq x, y \leq 1$,

The probability density functions are $f(x) = 1$ ($0 \leq x \leq 1$) and $g(y) = 1$ ($0 \leq y \leq 1$). The integration over x simply removes the delta function.

$$p(z) = \int_0^1 dy \int_0^1 \delta(z - x - y) dx$$

The range of the subsequent y -integration is clear from the line $y = -x + z$ drawn in the sample space of Fig. 1: when $0 \leq z \leq 1$, we have $0 \leq y \leq z$, and when $1 \leq z \leq 2$, we have $z - 1 \leq y \leq 1$.

$$p(z) = \begin{cases} \int_0^z dy = z & 0 \leq z \leq 1 \\ \int_{z-1}^1 dy = 2 - z & 1 \leq z \leq 2 \end{cases}$$

【Example 2】 For exponential random variables x and y ,
 $f_X(x) = \lambda e^{-\lambda x}$ ($x \geq 0$), $g_Y(y) = \lambda e^{-\lambda y}$ ($y \geq 0$),

$$\begin{aligned} f_G(z) &= \lambda^2 \int \delta(z - x - y) e^{-\lambda x} e^{-\lambda y} dx dy \\ &= \lambda^2 e^{-\lambda z} \int_0^z dy \quad (x = z - y \geq 0, y \geq 0 \rightarrow 0 \leq y \leq z) \\ &= \lambda^2 z e^{-\lambda z} \end{aligned}$$

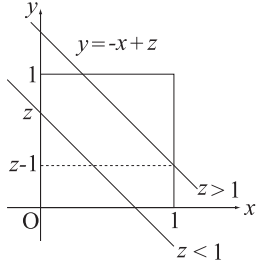


Fig. 1: Sample space and $y = -x + z$

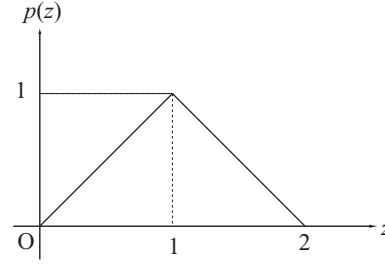


Fig. 2: Probability density function $p(z)$

The gamma distribution $\Gamma(\alpha, \lambda)$ is defined by

$$f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x}$$

The exponential distribution is the gamma distribution with $\alpha = 1$. The sum of mutually independent exponentially distributed random variables follows the gamma distribution with $\alpha = 2$.

【Example 3】 For standard normal random variables x and y ,

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad g(y) = \frac{1}{\sqrt{2\pi}} e^{-y^2/2},$$

$$\begin{aligned} f_N(z) &= \frac{1}{2\pi} \int \delta(z - x - y) e^{-(x^2+y^2)/2} dx dy \\ &= \frac{1}{2\pi} \int e^{-((z-y)^2+y^2)/2} dy \\ &= \frac{1}{2\pi} e^{-z^2/4} \int e^{-(y-z/2)^2} dy \end{aligned}$$

Let $t = y - z/2$. Then

$$\int e^{-(y-z/2)^2} dy = \int e^{-t^2} dt = \sqrt{\pi}$$

Therefore,

$$f_N(z) = \frac{1}{2\sqrt{\pi}} e^{-z^2/4}$$

This is the normal distribution with mean $\mu = 0$ and variance $\sigma^2 = 2$.

(b) $\mathbf{Z = XY}$

$$q(z) = \int \delta(z - xy) f(x) g(y) dx dy \quad (2)$$

To evaluate this integral, we use the following property of the δ function.

$$\delta(u(x)) = \frac{1}{|u'(x)|} \delta(x - \alpha) \quad (3)$$

Here, α is a root of $u(x) = 0$. In this case, we have $u(x) = z - xy$, and $u'(x) = -y$. Substituting equation (3) into equation (2), we obtain

$$q(z) = \int \frac{1}{|y|} f(z/y) g(y) dy$$

as desired.

【Example 1】 When the random variables x and y follow the uniform distribution on $0 \leq x, y \leq 1$,

$$q(z) = \int_z^1 \frac{dy}{y} = -\ln z$$

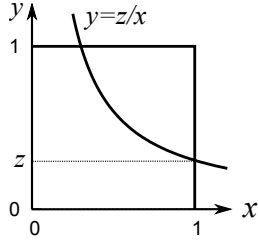


Fig. 3: Sample space and $y = z/x$

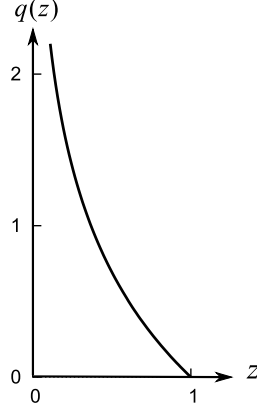


Fig. 4: Probability density function $q(z)$

【Example 2】 Example 2 When the random variables x and y follow the exponential distribution,

$$\begin{aligned} f(z) &= \lambda^2 \int \delta(z - xy) e^{-\lambda x} e^{-\lambda y} dx dy \\ &= \lambda^2 \int_0^\infty \frac{1}{y} e^{-\lambda(z/y+y)} dy \end{aligned}$$

Let $t = \lambda y$. Then

$$f(z) = \lambda^2 \int_0^\infty \frac{1}{t} e^{-(\lambda^2 z/t+t)} dt \tag{4}$$

Now consider the function

$$K_n(z) = \frac{1}{2} \left(\frac{z}{2}\right)^n \int_0^\infty \exp\left(-t - \frac{z^2}{4t}\right) \cdot t^{-n-1} dt$$

which is called the modified Bessel function of the second kind. In particular, when $n = 0$,

$$K_0(z) = \frac{1}{2} \int_0^\infty \exp\left(-t - \frac{z^2}{4t}\right) \cdot t^{-1} dt \tag{5}$$

so, using equation (5), equation (4) can be rewritten as

$$f(z) = 2\lambda^2 K_0(2\lambda\sqrt{z})$$

which gives the result.

【Example 3】 When the random variables x and y follow the standard normal distribution ,

$$f_Z(z) = \frac{1}{2\pi} \int_{-\infty}^\infty dy \int_{-\infty}^\infty \delta(z - xy) e^{-(x^2+y^2)/2} dx$$

$$\begin{aligned}
&= \frac{1}{2\pi} \int_{-\infty}^{\infty} \frac{1}{|y|} e^{-(z^2/y^2+y^2)/2} dy \\
&= \frac{1}{2\pi} \left\{ - \int_{-\infty}^0 \frac{1}{y} e^{-(z^2/y^2+y^2)/2} dy + \int_0^{\infty} \frac{1}{y} e^{-(z^2/y^2+y^2)/2} dy \right\} \\
&= \frac{1}{\pi} \int_0^{\infty} \frac{1}{y} e^{-(z^2/y^2+y^2)/2} dy
\end{aligned}$$

Let $t = \frac{y^2}{2}$. Since $\frac{dy}{y} = \frac{dt}{2t}$,

$$\begin{aligned}
f_Z(z) &= \frac{1}{2\pi} \int_0^{\infty} \frac{e^{-t-z^2/4t}}{t} dt \\
&= \frac{1}{\pi} K_0(z)
\end{aligned}$$

Thus, the modified Bessel function of the second kind appears once again.

(c) $Z = X/Y$

$$\begin{aligned}
r(z) &= \int \delta(z - x/y) f(x)g(y) dx dy \\
&= \int |y| f(zy)g(y) dy
\end{aligned}$$

Here, $u(x) = z - x/y$ and $u'(x) = -1/y$ in equation (3) which is used for the integration over x .

【Example 1】 When the random variables x and y follow the uniform distribution on $0 \leq x, y \leq 1$,

$$r(z) = \int_0^1 dy \int_0^1 \delta(z - x/y) dx$$

The range of the subsequent y -integration is clear from the line $y = x/z$ drawn in the sample space of Fig. (5): when $0 \leq z \leq 1$, we have $0 \leq y \leq 1$, and when $1 \leq z$, we have $0 \leq y \leq 1/z$.

$$r(z) = \begin{cases} \int_0^1 y dy = \left[\frac{y^2}{2} \right]_0^1 = \frac{1}{2} & 0 \leq z \leq 1 \\ \int_0^{1/z} y dy = \left[\frac{y^2}{2} \right]_0^{1/z} = \frac{1}{2z^2} & 1 \leq z \end{cases}$$

【Example 2】 When the random variables x and y follow the exponential distribution,

$$\begin{aligned}
f(z) &= \lambda^2 \int \delta(z - x/y) e^{-\lambda x} e^{-\lambda y} dx dy \\
&= \lambda^2 \int_0^{\infty} y e^{-\lambda(z+1)y} dy
\end{aligned}$$

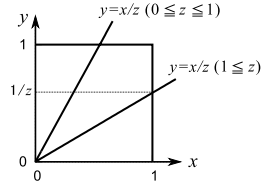


Fig. 5: Sample space and $y = x/z$

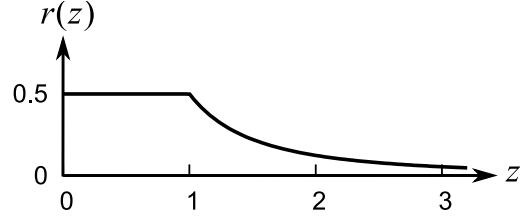


Fig. 6: Probability density function $r(z)$

Let $t = \lambda(z + 1)y$. Then

$$\begin{aligned}
 f(z) &= \frac{1}{(z+1)^2} \int_0^\infty t e^{-t} dt \\
 &= \frac{1}{(z+1)^2} \int_0^\infty t (-e^{-t})' dt \\
 &= \frac{1}{(z+1)^2} \left\{ [-t e^{-t}]_0^\infty + \int_0^\infty e^{-t} dt \right\} \\
 &= \frac{1}{(z+1)^2}
 \end{aligned}$$

That is,

$$f(z) = \frac{1}{(z+1)^2}$$

as desired.

【Example 3】 For standard normal random variables x and y ,

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad g(y) = \frac{1}{\sqrt{2\pi}} e^{-y^2/2},$$

$$\begin{aligned}
 f_C(z) &= \frac{1}{2\pi} \int_{-\infty}^\infty dy \int_{-\infty}^\infty \delta(z - x/y) e^{-(x^2+y^2)/2} dx \\
 &= \frac{1}{2\pi} \int_{-\infty}^\infty |y| e^{-(z^2+1)y^2/2} dy \\
 &= \frac{1}{2\pi} \left\{ - \int_{-\infty}^0 y e^{-(z^2+1)y^2/2} dy + \int_0^\infty y e^{-(z^2+1)y^2/2} dy \right\} \\
 &= \frac{1}{\pi} \int_0^\infty y e^{-(z^2+1)y^2/2} dy
 \end{aligned}$$

Using $t = (1 + z^2)y^2/2$ in place of the integration variable y ,

$$\int_0^\infty y e^{-(z^2+1)y^2/2} dy = \frac{1}{1+z^2} \int_0^\infty e^{-t} dt = \frac{1}{1+z^2}$$

Therefore,

$$f_C(z) = \frac{1}{\pi(1+z^2)}$$

This is called the Cauchy distribution.

(d) $Z = X^Y$

$$\begin{aligned} s(z) &= \int \delta(z - x^y) f(x) g(y) dx dy \\ &= \int \frac{1}{|y z^{1-1/y}|} f(z^{1/y}) g(y) dy \end{aligned}$$

Here, $u(x) = z - x^y$ and $u'(x) = -yx^{y-1}$ in equation (3) which is used for the integration over x .

【Example】 Example For uniform random variables x and y on $0 \leq x, y \leq 1$,

$$\begin{aligned} s(z) &= \int_0^1 dy \int_0^1 \delta(z - x^y) dx \\ &= \int_0^1 \frac{1}{y z^{1-1/y}} dy \\ &= \frac{1}{z} \int_0^1 \frac{z^{1/y}}{y} dy \\ &= \frac{1}{z} \int_0^1 \frac{e^{(\ln z)/y}}{y} dy \end{aligned}$$

Let $t = \frac{\ln z}{y}$. Then

$$e^{(\ln z)/y} = e^t, \quad dt = -\ln z \frac{dy}{y^2} = -t \frac{dy}{y}$$

Therefore,

$$s(z) = -\frac{1}{z} \int_{-\infty}^{\ln z} \frac{e^t}{t} dt$$

which gives the result.

The integral appearing here is called the exponential integral,

$$Ei(\ln z) = \int_{-\infty}^{\ln z} \frac{e^t}{t} dt$$

so it can be written as follows, and hence

$$s(z) = -\frac{1}{z} Ei(\ln z), \quad 0 \leq z \leq 1$$

as desired.

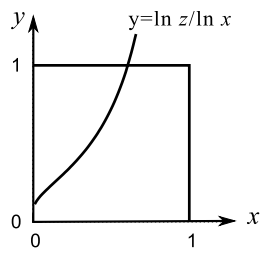


Fig. 7: Sample space and $y = \ln z / \ln x$

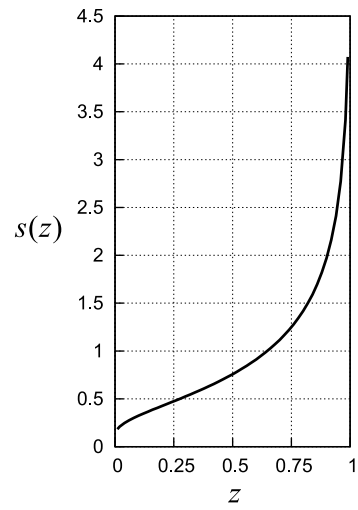


Fig. 8: Probability density function $s(z)$